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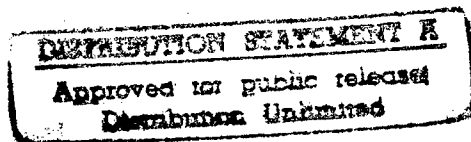
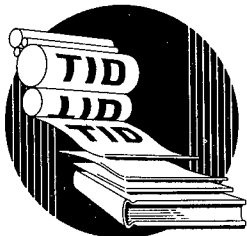
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ON THE STATISTICS OF LUMINESCENT  
COUNTER SYSTEMS

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## ON THE STATISTICS OF LUMINESCENT COUNTER SYSTEMS

By Frederick Seitz and D. W. Mueller

### 1. INTRODUCTION

The type of crystal counter which depends upon the combination of luminescent crystals and a photomultiplier tube shows promise of being of great service in the detection of radiations both because of its high sensitivity and speed of registry and recovery. This device has been developed by a large number of individuals, almost too numerous to mention; however, the origin of the system appears to rest with Coltman and Marshall,<sup>1</sup> who employed powdered luminescent materials of the type used in previous commercial luminescent systems, and with Broser and Kallmann,<sup>2</sup> who first appreciated the advantages of employing large, transparent, luminescent crystals and introduced organic materials.

The purpose of this paper is to analyze some of the factors which influence the statistical behavior of luminescent counter systems, in order to evaluate the limits within which a counter may be used in making a particular type of measurement. The problems of interest range over a wide spectrum of possibilities. However, the problem on which attention is focused for immediate purposes in order to provide a practical objective is the following:

A crystal-counter system is employed to count the gamma rays emitted from a source in time  $T$ . If  $N$  gamma rays are emitted, what is the most probable number that will be counted and what is the range of variation to be expected? An attempt is made to examine this problem in a sufficiently general way that the results will have value for a much broader group of problems.

It is interesting to consider the component parts of this problem in order to be able to examine the sources of statistical variations. The components are as follows:

1. The source, even if constant in the sense that it remains unchanged during the time  $T$ , will contribute to the statistical variation since the gamma rays are usually emitted at random. For simplicity, it is assumed that the time  $T$  is sufficiently short that variations in the source strength can be neglected and that the statistical variations in emission of gamma rays can be treated on the basis of a Poisson distribution.

2. Unless the source is completely surrounded by the luminescent material, some of the gamma rays will not pass through this material and hence will certainly fail to be registered. The average fraction which passed through the material is designated by  $f$ , so that the average number of gamma rays which pass through the detecting system, if  $N$  are emitted from the source, is

$$\nu = fN \quad (1)$$

If the source is isotropic,  $f$  will be determined simply by the solid angle subtended by the crystal system; otherwise a somewhat more involved calculation is needed to determine  $f$ .

3. A given gamma ray may or may not produce an ionizing pulse within the luminescent crystal. The possible mechanisms for producing such a pulse are the photoelectric effect, the Compton effect, and pair production. In the first and third cases the gamma ray transmits all its energy to the crystal provided the energetic electrons produced by the gamma ray do not escape from the crystal. A greater statistical variation is possible when the range of gamma-ray energy and the atomic number are such that the Compton effect predominates. This would be the case, for example, if the luminescent material were one of the organic types such as naphthalene or anthracene and if the gamma rays had an energy in the neighborhood of 2 Mev.

The probability of a Compton encounter may be described in terms of the mean free path  $\lambda$  for the process, namely

$$\lambda = 1/n_e \sigma_c \quad (2)$$

where  $n_e$  is the density of electrons in the luminescent material and  $\sigma_c$  is the Compton cross section per electron. If  $d$  is the thickness of the luminescent material in the direction in which the incident gamma rays are traveling, the probability that a given gamma ray will pass through the system without producing a Compton electron is  $e^{-\gamma}$  where

$$\alpha = d/\lambda \quad (3)$$

The initial energy  $k_0$  of the gamma ray and the energy  $k$  after the collision are related by the equation

$$\frac{k}{k_0} = \frac{1}{1 + \gamma(1 - \cos \theta)} \quad (4)$$

where  $\theta$  is the angle between the incident and scattered quantum and  $\gamma$  is the energy of the incident gamma ray expressed in units of the rest mass of the electron (507 kev). The energy gained by the electron is  $\epsilon = k_0 - k$ . From Eq. 4 the relation

$$d(\cos \theta) = \frac{k_0}{\gamma k^2} dk \quad (5)$$

may be readily derived, connecting the differentials of  $\cos \theta$  and  $k$ . The differential cross section  $d\phi$  for scattering into solid angle  $d\Omega$  is

$$d\phi = \frac{r_0^2}{2} \frac{d\Omega}{k_0^2} \left( \frac{k_0}{k} + \frac{k}{k_0} - \sin^2 \theta \right) \quad (6)$$

in which  $r_0$  is the classical electron radius  $e^2/mc^2$ . If the relation  $d\Omega = 2\pi \sin \theta d\theta$  is used and  $d\theta$  is replaced by  $dk$  with the use of Eq. 5 the following equation is obtained.

$$d\phi = \pi r_0^2 \frac{1}{k_0 \gamma} \left( \frac{k_0}{k} + \frac{k}{k_0} - \sin^2 \theta \right) dk \quad (7)$$

This relation is to be employed in the range of  $k$  varying from  $k_0$  to  $k_0/(1 + 2\gamma)$  corresponding to the range of  $\theta$  from 0 to  $\pi$ . The quantity in parenthesis in Eq. 7 has the following values when  $\theta$  takes the values 0,  $\pi/2$ , and  $\pi$ :

$$\begin{aligned} \theta = 0 & \quad 2 \\ \theta = \pi/2 & \quad \frac{1 + \gamma + \gamma^2}{1 + \gamma} \\ \theta = \pi & \quad 2 \frac{1 + 2\gamma + 2\gamma^2}{1 + 2\gamma} \end{aligned} \quad (8)$$

For values of  $\gamma$  not larger than about 4, this variation is sufficiently small that it is reasonably good to assume that  $k$  has equal probability of falling in any part of the allowed range, or that the knocked-on electron has equal probabilities of receiving any energy in the range from 0 to  $2\gamma/(1 + 2\gamma)$  in units of  $k_0$ . For very large values of  $\gamma$ , the  $\sin^2 \theta$  term in parenthesis in Eq. 7 may be neglected for the most interesting collisions. It is then clear from the remaining terms in parenthesis that collisions in which  $k$  is small compared with  $k_0$  are preferred over those in which  $k$  is near  $k_0$ .

The degree of preference is not exceedingly great for values of  $\gamma$  in the normal radioactive range, and it was assumed that the probability per unit energy range is constant within the allowed limits. This gives the maximum statistical variation to be expected in a given Compton process.

The gamma ray may conceivably make a number of Compton encounters in passing through the crystal. There are two interesting extreme cases to consider which are referred to as the "thin" and "thick" approximations. In the thin case, which corresponds to values of  $\alpha$  appreciably less than unity, the gamma ray has much smaller probability of making two collisions than one collision. In this case it is assumed that  $\alpha$  may be chosen to be a constant for each successive collision as if its energy were not greatly affected by successive Compton encounters. In this event the probability that the gamma ray will make  $n$  encounters may be described by the distribution function

$$P_n = \frac{\alpha^n e^{-\alpha}}{n!} \quad (9)$$

for the range of  $n$  of practical interest.

In the thick approximation the gamma ray transfers all its energy to the luminescent material in a succession of encounters once it has made the first encounter. Thus this case is equivalent to that in which the gamma ray transfers its energy by means of the photoelectric effect or pair production, provided the electrons produced do not escape. These last two cases differ from the thick approximation only with respect to the geometrical distribution of points within the crystal at which the electrons are released—a difference which is not considered here.

The thin approximation is best achieved by employing a very thin crystal so that  $\alpha$  is small compared with unity and also employing soft gamma rays, for which  $\gamma$  is 1 or less, which lose relatively little energy in a Compton encounter. It is probably not a case which would be met in practice but is interesting as one statistical extreme. It should be remarked that this limit cannot be achieved by going to very soft radiation, for such radiation is scattered almost isotropically. The distance which the scattered photon must traverse is usually different from that which the original photon would have had to travel to pass through the crystal because it is traveling in a different direction. Thus  $\alpha$  is not a constant in this limit even though the energy of the photon is not greatly altered by a Compton collision. The thick case can evidently be achieved by using a thick crystal and is the one that will be met more commonly in practice.

4. The number of luminescent quanta which the crystal emits can vary even when the energy transmitted to the crystal is fixed because of statistical fluctuations in the manner in which the exciting radiation is distributed among the different excited states of the medium. This type of statistical fluctuation is partly responsible for the straggling in range of heavy ionizing particles as they pass through matter. In order of magnitude the fractional variation in the number of light quanta is  $1/\sqrt{\eta}$ , where  $\eta$  is the average number. Since in this paper cases in which  $\eta$  is 1000 or larger are of interest, corresponding to Compton encounters in which the knocked-on electron gains several hundred kev of energy, this source of statistical variation will be neglected. It could be significant in cases in which the particle being detected produces very few light quanta, as for very soft beta rays or x rays.

5. Only a fraction of the light quanta produced in the luminescent crystal will reach the photoelectric surface. The fraction  $Q$  which does is determined primarily by geometrical factors involving the angular distribution of emitted light and the angle subtended by the photosurface relative to the luminescent material.  $Q$  will be in the neighborhood of 0.5 for a relatively thin layer of luminescent material which is immediately adjacent to the photosurface but may be considerably smaller if the luminescent crystal is somewhat farther away. It may be enhanced by placing a reflecting backing on the luminescent material or by employing other devices which cause the light to be "funneled" toward the photocathode.

6. Only a fraction  $p$  of the photons striking the photosurface will eject electrons from it. This parameter appears<sup>4</sup> to be about 0.03 for the type of photosurface in which the photons penetrate the photosurface and electrons are ejected from the back side, and it appears to be about 0.05 for the type of photosurface for which electrons are ejected from the front surface. (Morton and Mitchell<sup>4</sup> have shown that the pulse-height distribution is broader than that expected on the basis of a Poisson distribution of electrons at each stage.)

7. The electrons ejected from the photocathode will give rise to pulses of various size, depending upon the accidents which befall the primary photoelectron and the secondaries which it emits from the multiplying surfaces. Actually there are two problems associated with an analysis of the pulse distribution: first, the problem of determining the probability that the photoelectron will actually create a measurable pulse and second, the problem of determining distribution of pulse sizes when pulses are generated. If the secondary emission ratio is  $s$ , the probability that the photoelectron will not eject a secondary from the first stage of the multiplier is  $e^{-s}$ , provided it is assumed that the emission of secondaries is random. This probability is of the order of a few per cent for normal values of  $s$  (between 3 and 5) and is essentially equal to the probability that the primary electron will not generate a pulse. Since the percentage of uncertainty in  $p$  is at least as large as this, this factor may be combined with  $p$  in the following discussion, and the assumption may be made that a measurable pulse is produced whenever an electron is ejected. The distribution of pulse sizes has been measured by Engstrom<sup>4</sup> using a light source. Later, his results will be approximated with an appropriate mathematical function. Evidently the pulse distribution is not important if the luminescent counter is employed simply as a counter of events and if a pulse of arbitrary size can be employed as the signal for a significant count. Knowledge of the distribution becomes important, however, if a pulse discriminator is employed so that only pulses larger than a certain size are counted (as when a noise background is eliminated) or if the integrated current of the photomultiplier is recorded. The first of these cases may be treated by redefining the parameter  $p$  as the probability that an observable pulse is measured when a photon strikes the cathode and introducing measured values of this quantity. The second case is discussed in detail.

## 2. THE GENERATING FUNCTION AND ITS APPLICATIONS TO THE PROBLEM<sup>5</sup>

The generating function was introduced into probability theory very early in its development, and some of its properties are described in textbooks, for example, the book by Uspensky.<sup>5</sup> The authors have benefitted by reading a mimeographed survey of the subject by O. R. Frisch. An account of some of the relations employed here is given by Jorgenson.<sup>5</sup>

The aggregate contribution of the various unit parts of the photomultiplier system to the statistical variation of the system can be determined most simply with the use of generating functions appropriate to each stage. If  $p_n$  is the probability that a given observation shall yield  $n$  events, for example, that the source in the problem will emit  $n$  gamma rays in time  $T$ , the generating function  $G(\epsilon)$  for the process of observation is defined by the series

$$G(\epsilon) = p_0\epsilon^0 + p_1\epsilon^1 + p_2\epsilon^2 + \dots + p_n\epsilon^n + \dots \quad (10)$$

The generating function is readily found to possess the following properties

$$G(0) = p_0 \quad G(1) = 1 \quad (11)$$

The mean value  $m$  of a series of observations, namely

$$m = \sum_n np_n \quad (12)$$

is readily seen to satisfy the relation

$$m = \left( \frac{dG}{d\epsilon} \right)_{\epsilon=1} \quad (13)$$

Similarly the variance of a sequence of observations, defined by the relation

$$v = \sum_n (n - m)^2 p_n = \sum_n n^2 p_n - m^2 \quad (14)$$

is readily found to be related to the generating function by the equation

$$v = \left[ \frac{d^2 G}{d\epsilon^2} + \frac{dG}{d\epsilon} - \left( \frac{dG}{d\epsilon} \right)^2 \right]_{\epsilon=1} = \left( \frac{d^2 G}{d\epsilon^2} \right)_{\epsilon=1} + m - m^2 \quad (15)$$

In the case of the Poisson distribution

$$p_n = \frac{\alpha^n}{n!} e^{-\alpha} \quad (16)$$

$G(\epsilon)$  is readily found to be

$$G(\epsilon) = e^{\alpha(\epsilon-1)} \quad (17)$$

whence

$$m = \alpha \quad v = \alpha \quad (18)$$

In the following calculations the ratio  $\sqrt{v}/m$  is employed to provide a measure of the fractional deviation from the mean or the fractional deviation. In the case of the Poisson distribution this quantity is the familiar ratio  $1/\sqrt{\alpha}$ .

Although the generating function is interesting and useful because of the properties already outlined, its real service appears when the following two additional properties are considered:

I. Suppose that, instead of making one observation of the number of events of interest (such as the number of gamma rays emitted from the source in time  $T$ ), two observations are made (e.g., for two time intervals  $T$ ) and ask for the probability that  $n$  events are observed en toto is asked for. The probability for this is the sum

$$p_n p_0 + p_{n-1} p_1 + p_{n-2} p_2 + \dots + p_0 p_n$$

which is the coefficient of  $\epsilon^n$  in the expansion of  $G^2(\epsilon)$ . This is a special case of the more general theorem: The generating function governing the probability distribution of the sum of  $r$  identical observations is  $G^r(\epsilon)$  if  $G(\epsilon)$  is the generating function for a single observation.

II. Suppose next that a situation is being dealt with in which each member of a set of initial events that are statistically distributed (e.g., gamma rays from a source) can give rise to a series of events of possibly different type (e.g., production of Compton electrons) and the statistical distribution of the second type of event is asked for. Let  $G_1(\epsilon)$  be the generating function for the first type of event (e.g., the number of gamma rays emitted by the source in a given time for the example under consideration) and  $G_2(\epsilon)$  be the generating function for the number of events of the second type associated with one primary event (e.g., the number of Compton recoils produced by a single gamma ray). It is readily shown that the generating function  $G_{II}(\epsilon)$  for the number of events of the second type when the statistical variation of the number of events of the first type is taken into account is given by

$$G_{II}(\epsilon) = G_1[G_2(\epsilon)] \quad (19)$$

The validity of this theorem may readily be demonstrated by writing  $G_{II}$  in the form

$$G_{II}(\epsilon) = p_0 G_2^0(\epsilon) + p_1 G_2^1(\epsilon) + p_2 G_2^2(\epsilon) + p_3 G_2^3(\epsilon) + \dots + p_n G_2^n(\epsilon) + \dots \quad (20)$$

in which  $p_n$  is the probability of  $n$  events of the first type, so that

$$G_1(\epsilon) = p_0 \epsilon^0 + p_1 \epsilon^1 + p_2 \epsilon^2 + \dots + p_n \epsilon^n + \dots$$

The coefficient of  $p_n$  in Eq. 20 is the generating function for the total number of events of the second type when it is known that  $n$  events of the first type have occurred, in accordance with theorem I. This coefficient appears suitably weighted with the probability that  $n$  primary events will occur.

Using Eqs. 13 and 14, the mean and variance associated with the generating functions

$$G_I(\epsilon) = G^r(\epsilon) \quad G_{II}(\epsilon) = G_1[G_2(\epsilon)] \quad (21)$$

may readily be found which occur in the cases I and II described above. The results are, respectively

$$m_I = rm \quad v_I = rv \quad (22)$$

in which  $m$  and  $v$  are the mean and variance associated with a single observation in case I, and

$$m_{II} = m_1 m_2 \quad v_{II} = v_1 m_2^2 + v_2 m_1 \quad (23)$$

It is clear that, if case II were extended to that in which the second type of event can give rise to a third type (e.g., if a Compton electron can give rise to ion pairs or to luminescent quanta) which is statistically distributed in accordance with a generating function  $G_3(\epsilon)$ , the complete generating function which takes account of the statistical variation in events of the three types is

$$G_{III} = G_1[G_2(\epsilon)] \quad (24)$$

for which the mean and variance are, by analogy with Eq. 23

$$m_{III} = m_{II} m_3 \quad v_{III} = v_{II} m_3^2 + v_3 m_{II} \quad (25)$$

The appropriate form of generating functions to be employed in each of the constituent processes described in Sec. 1 will now be examined.

1. Emission from source. Since the gamma rays are emitted at random, the appropriate generating function is of the Poisson type, Eq. 17, namely

$$G_1(\epsilon) = e^{N(\epsilon-1)} \quad (26)$$

in which  $N$  is the average number of gamma rays emitted in the time  $T$ .

2. Passage of gamma rays into system. A given gamma ray either does or does not enter the crystal. If the probability that it does is  $f$ , the generating function for this event is simply

$$G_2(\epsilon) = (1 - f) + f\epsilon \quad (27)$$

Using Eq. 19, it is readily found that the generating function  $G_2'(\epsilon)$ , giving the distribution of probabilities that the gamma rays emitted at random by the source enter the crystal, is

$$G_2'(\epsilon) = G_1[G_2(\epsilon)] = e^{Nf(\epsilon-1)} = e^{\nu(\epsilon-1)} \quad (28)$$

where  $\nu = fN$ .

3. Generation of photons. As stated in the introduction, the assumption is that a fixed fraction of the energy which the gamma photon gives up to the crystal is transformed into light quanta. If this energy is  $E$ , the number of light quanta produced is then

$$\eta = \beta E \quad (29)$$

where  $\beta$  is a factor measuring the efficiency with which the luminescent crystal converts the excitation energy it receives into light quanta. If  $h\nu$  is the average energy of the luminescent quanta emitted,  $\beta h\nu$  is the fraction of the energy of excitation which appears in the form of luminescent radiation. This may be as large as 0.20 for some of the most efficient materials but can easily be much smaller. According to Broser, Kallman, and Martius<sup>6</sup> the efficiencies of energy conversion in zinc sulfide activated with silver and in the organic materials naphthalene, diphenyl, and phenanthrene are given in Table 1.

Table 1—Efficiency of Energy Conversion in Luminescent Materials under Gamma-ray Excitation

(After Broser, Kallmann, and Martius. Values in fractions.)

	$\beta h\nu$
ZnS:Ag	0.135
Naphthalene	0.05
Diphenyl	0.075
Phenanthrene	0.11



The investigators find somewhat different efficiencies for beta-ray excitation. Similarly, Colman, Ebbighausen, and Altar<sup>7</sup> have found the energy conversion in calcium tungstate to be 5.0 per cent for x rays. The interest here is in detailed values of the efficiency in Sec. 3.

As mentioned in the introduction there are two extreme approximations that are of interest, namely, those designated as "thin" and "thick." In the second case, all the energy of the gamma ray is transmitted to the crystal once a first collision has occurred. The number of quanta emitted is then equal to  $\eta_0$ , the value of  $\eta$  when  $k_0$  is the energy of the gamma ray. If  $c$  is the probability that such a collision occurs, the generating function for the number of quanta is evidently

$$S_3(\epsilon) = (1 - c) + c\epsilon\eta_0 \quad (30)$$

If this is combined with Eq. 28, the complete generating function for the production of luminescent quanta in the thick case is

$$S'_3 = \exp [\nu c(\epsilon\eta_0 - 1)] \quad (31)$$

In the thin case there are two sources of statistical variation, for both the number of Compton encounters and the energy transferred to the counter per collision may vary. The first of these quantities is distributed in accordance with the Poisson law, Eq. 17, in the ideal thin case, for which the generating function is

$$H_3 = e^{\alpha(\epsilon-1)}$$

where  $\alpha$  is the ratio (Eq. 3). The energy which the Compton electron receives is randomly distributed between 0 and the maximum value  $2k_0\gamma/(1 + 2\gamma)$  in the approximation described in paragraph 3 of the introduction. This means that the number of quanta generated will vary between 0 and a maximum  $\eta_m$ , where

$$\eta_m = \beta \frac{2k_0\gamma}{1 + 2\gamma} \quad (32)$$

in which  $\beta$  is the efficiency factor appearing in Eq. 29. A generating function for this random distribution is readily constructed by treating  $\eta$  as a continuous variable and is

$$K_3(\epsilon) = \frac{1}{\eta_m} \int_0^{\eta_m} \epsilon^\eta d\eta = \frac{\exp(\eta_m \log \epsilon) - 1}{\eta_m \log \epsilon} \quad (33)$$

for which the mean and variance are  $\eta_m/2$  and  $\eta_m^2/12$ , respectively. The complete generating function for the number of quanta associated with a single gamma ray is  $H_3[K_3(\epsilon)]$ .

4. Emission of photoelectrons. A given light quantum either does or does not emit a photoelectron from the photosurface of the multiplier. The probability that it does is  $\Omega p$ , so that the generating function for this process is

$$G_4(\epsilon) = [(1 - \Omega p) + \Omega p\epsilon] \quad (34)$$

As stated in paragraph 7 of the introduction, it is assumed that a measurable pulse is associated with each photoelectron ejected from the cathode of the multiplier.

5. Generating function for pulse distribution. Engstrom<sup>4</sup> has measured the pulse-height distribution of a typical multiplier tube. His empirical distribution is represented by the analytical function

$$f(h) = Ah^2 \exp(-h/\rho) \quad (35)$$

in which  $h$  is the pulse height on an arbitrary scale,  $\rho$  is a constant measuring the width of the distribution, and  $A$  is a normalization factor  $1/2\rho^3$ . A generating function

$$G_5(\epsilon) = 1/(1 - \rho \log \epsilon)^3 \quad (36)$$

may readily be constructed for this distribution. The mean and variance are

$$m_s = 3\rho \quad v = 3\rho^2 \quad (37)$$

It will be seen later in this paper that it is not necessary to know  $\rho$  in order to determine the fractional deviation of interest here.

### 3. CORRELATION BETWEEN GAMMA RAYS FROM SOURCE AND PULSES IN MULTIPLIERS

The probability that a gamma ray from the source will produce a pulse in the multiplier is discussed in this section. The thick and thin cases are discussed separately.

#### A. Thick Case

In the thick case, a gamma ray passing through the crystal has probability  $c$  of making an encounter, in which case it generates  $\eta_0$  photons. The distribution of such encounters is random, being governed by the Poisson distribution. The average number is  $Nfc = \nu c$ , and the deviation is  $\nu c$ . Thus, as far as luminescent pulses are concerned, the effective strength of the source is  $\nu c$ .

The probability that  $n$  of the  $\eta_0$  light quanta will eject photoelectrons from the multiplier is given by the generating function

$$[G_4(\epsilon)]^{\eta_0} = [(1 - \Omega p) + \Omega p \epsilon]^{\eta_0} \quad (38)$$

The probability that none will eject electrons is  $(1 - \Omega p)^{\eta_0}$ , so that the probability of observing a pulse, if one electron is sufficient to produce an observable pulse, is

$$P = 1 - (1 - \Omega p)^{\eta_0} \quad (39)$$

Since  $\eta_0$  is usually large compared with unity, this may be approximated by

$$P = 1 - \exp(-\eta_0 \Omega p) \quad (40)$$

in which the quantity

$$\eta_0 \Omega p \quad (41)$$

is the average number of photoelectrons emitted from the cathode. With the use of the rule, Eq. 25, for determining the mean and variance of a chain of events, the mean number of counts is

$$M = Nfc[1 - \exp(-\eta_0 \Omega p)] = NfcP \quad (42)$$

whereas the variance is

$$V = NfcP \quad (43)$$

The fractional variance is

$$\frac{\sqrt{V}}{M} = \left( \frac{1}{NfcP} \right)^{1/2} \quad (44)$$

#### B. Thin Case

There are four statistical processes in the chain extending from the passage of gamma rays into the crystal to the ejection of electrons from the photocathode, namely, those described by the generating functions  $G'_2$ ,  $H_3$ ,  $K_3$ , and  $G_4$  of the preceding section. The number of electrons ejected from the cathode when a single gamma ray passes through the crystal is governed by the generating function

$$E(\epsilon) = H_3 \left\{ K_3 [G_4(\epsilon)] \right\} \quad (45)$$

The probability that no electron will be ejected and hence that no pulse will be recorded is  $E(0)$ , so that the probability of a pulse is  $1 - E(0)$ , and the generating function for pulses is

$$\{E(0) + [1 - E(0)]\epsilon\} \quad (46)$$

Hence the mean and variance in the number of pulses are

$$M = V = Nf[1 - E(0)] \quad (47)$$

Since  $\Omega p$  is of the order of 3 per cent even when  $\Omega$  is unity, it is readily found that  $K_3[G_4(0)]$  can be approximated by the expression

$$K_3[G_4(0)] = \frac{[1 - \exp(-\eta_m \Omega p)]}{\eta_m \Omega p} \quad (48)$$

A simple examination of  $E(0)$  shows that it approaches  $e^{-\alpha}$  when  $\eta_m \Omega p$  is large compared with unity and approaches  $\exp(-\alpha \eta_m \Omega p / 2)$  when  $\eta_m \Omega p$  is small. It may be concluded that, in both the thick and thin cases, the counts are governed by a Poisson distribution and that it is desirable to have the quantities  $c$  and  $1 - e^{-\alpha}$  as near unity as possible and the quantities  $\eta_0 \Omega p$  and  $\eta_m \Omega p$  somewhat larger than unity, although there probably is little advantage to having them as large as 10.

Suppose one is dealing with gamma rays in the vicinity of 1.5 Mev, to provide a concrete example. In this case the mean free path for the Compton effect in a material such as naphthalene is of the order of 15 cm. Hence if the crystal is a cube 5 cm on an edge, the factor  $e^{-\alpha}$  is 0.72. The Compton electron will have an average energy of the order of 0.7 Mev, so that the average number of luminescent quanta produced is 10,000 if the energy efficiency is taken to be 0.05. Choosing  $p$  to be 0.03, it is found that  $\eta_m \Omega p / 2$  is  $300 \Omega$ . Hence  $\Omega$  should be at least  $10^{-2}$  if each Compton electron is expected to register with reasonable faithfulness. If it is assumed that the photosurface of the multiplier has an active area of about  $15 \text{ cm}^2$  and that this surface is 5 cm from the center of the crystal, the factor  $\Omega$  should be as large as 0.55 even if the photons are isotropically distributed, which would guarantee faithful counting of Compton encounters. The same photosurface would be more nearly borderline if the crystal were chosen to be a 10-cm cube and the surface were placed 10 cm from its center, for then  $\Omega$  would be about 0.01, which is very close to the limit set above. In fact, those Compton encounters which take place at points within the crystal which are most distant from the surface may fail to register if the photon distribution is isotropic. In this event it may prove profitable to employ a method of light funneling, for example, by covering all surfaces of the crystal except that opposite the multiplier with a reflecting metallic covering.

#### 4. PHOTOMULTIPLIER CURRENT

Consider next the current in the photomultiplier, or rather the charge which arrives at the anode end when  $N$  gamma rays are emitted from the source. The "thick" and "thin" cases are discussed separately once again.

##### A. Thick Case

In this case the distribution of charge in the photomultiplier may be regarded as if compounded of the three statistical processes which are described by the generating functions  $S'_3$ ,  $G_4$ , and  $G_5$  of Sec. 2. The first of these functions gives the distribution of photons in the crystal associated with the  $N$  gamma rays, the second function gives the distribution of the photoelectrons from the cathode of the multiplier, and the third function gives the distribution of pulses in the multiplier. The mean and variance of these distributions are as follows:

	Mean	Variance
$S'_3$	$\nu c \eta_0$	$\nu c \eta_0^2$
$G_4$	$\Omega p$	$\Omega p - (\Omega p)^2 \cong \Omega p$
$G_5$	$3p$	$3p^2$

The quantity  $(Qp)^2$  may be neglected in comparison with  $Qp$  since the latter is at most a few per cent.

By compounding these statistical quantities in accordance with the rule, Eq. 25, the following values of the mean and variance for the charge in the multiplier are obtained:

$$\begin{aligned} M &= 3\rho\nu c \eta_0 Qp \\ V &= 3\rho^2 \nu c \eta_0 Qp (4 + 3\eta_0 Qp) \end{aligned} \quad (49)$$

The fractional variance is

$$\frac{\sqrt{V}}{M} = \left[ \frac{4 + 3\eta_0 Qp}{3\nu c \eta_0 Qp} \right]^{1/2} \quad (50)$$

This quantity is independent of  $\rho$ , as pointed out previously. Moreover, it becomes independent of the quantity  $x = \eta_0 Qp$  when this quantity is large compared with unity. The condition placed upon  $x$  for this limit to be valid is somewhat more stringent than the condition required for faithful counting of luminescent pulses. That is,  $x$  must be larger than 5 for this approximation to be precise.

#### B. Thin Case

In this approximation the distribution of pulses is governed by a generating function that is compounded of the generating functions  $G'_2$ ,  $H_3$ ,  $K_3$ ,  $G_4$ , and  $G_5$ . Respectively, these correspond to the distribution of gamma rays in the crystal, the distribution of Compton encounters, the distribution of luminescent quanta produced in the crystal, the distribution of photoelectrons from the cathode, and the distribution of pulses in the multiplier. The corresponding means and variances are as follow:

	Mean	Variance
$G'_2$	$\nu$	$\nu$
$H_3$	$\alpha$	$\alpha$
$K_3$	$\eta_m/2$	$\eta_m^2/12$
$G_4$	$Qp$	$Qp$
$G_5$	$3\rho$	$3\rho^2$

The mean and variance for the distribution of pulses is found to be

$$\begin{aligned} M &= \frac{3}{2} \rho \nu \alpha \eta_m Qp \\ V &= \frac{3}{2} \rho^2 \nu \alpha \eta_m Qp \left[ 4 + \frac{1}{2} \eta_m Qp (4 + 3\alpha) \right] \end{aligned} \quad (51)$$

Once again it is noticed that  $\rho$  drops out of the fractional variance. Whenever the quantity  $y = \eta_m Qp$  is very small compared with unity, the fractional variance may be approximated by the expression

$$\frac{\sqrt{V}}{M} = \left( \frac{8/3}{\nu \alpha \eta_m Qp} \right)^{1/2} \quad (52)$$

In the opposite extreme, in which  $y$  is very large compared with unity, the fractional variance is

$$\frac{\sqrt{V}}{M} = \left( \frac{4 + 3\alpha}{3\nu\alpha} \right)^{1/2} \quad (53)$$

which approaches  $2/\sqrt{3\alpha\nu}$  if  $\alpha$  is small compared with unity and approaches  $1/\sqrt{\nu}$  if  $\alpha$  is very large. The latter case, in which  $\alpha$  is large, is in contradiction with the assumptions of the thin approximation; however, it is of mathematical interest.

## 5. FLUCTUATIONS IN CHARGE ON CONDENSOR

When dealing with a high-intensity source, it is frequently convenient to feed the current pulses from the photomultiplier into a condensor which is shunted with a high resistance and measure the voltage across the condensor in order to provide a measure of the average current which arrives at the condensor. This voltage exhibits fluctuations because the pulses are distributed statistically both in magnitude and in time. The influence of the distribution in time has been investigated by Schiff and Evans<sup>8</sup> for the case in which the pulses are equal in magnitude. The generalization of their results when the pulses vary in size is of interest here.

If the capacity of the condensor is  $C$  and the shunting resistance is  $R$ , the decay time for the shunted capacity is  $\tau = RC$ . A charge which is fed into the condensor at time  $t'$  will have decayed by a factor  $\exp [-(t - t')/\tau]$  by the later time  $t$ .

The assumption is that the charge associated with each pulse of the multiplier arrives in a time that is short compared with the decay time of the condensor. It is also assumed that the pulses are distributed in time in accordance with the distribution law governing the frequency with which gamma rays enter the luminescent crystal, that is, in accordance with the generating function  $G'_2(\epsilon)$  of Sec. 2 (see Eq. 28). Since interest is in specific intervals of time  $t$ ,  $\nu$  in Eq. 28 is replaced by  $nt$ , where  $n$  is the average number of gamma rays entering the crystal per unit time. Those gamma rays which do not excite the crystal will give rise to pulses of 0 size. For the purposes of this section, the generating function is designated for the pulse in the photomultiplier associated with the passage of a single gamma ray into the crystal by  $G(\epsilon)$ . The pulse size will be assumed to be expressed in units of charge.  $G(\epsilon)$  will differ in the soft and hard approximations but may be left arbitrary for the moment.

Consider the gamma rays which arrive in the time interval  $dt'$  between  $t'$  and  $t' + dt'$ . The generating function associated with the current they contribute to the condensor at the time  $t'$  is

$$1 + ndt' [G(\epsilon) - 1] \quad (54)$$

which is the expansion of  $G'_2[G(\epsilon)]$  in terms of  $dt'$  when  $\nu$  is replaced by  $ndt'$ . The mean value of the charge associated with this generating function is

$$ndt' G'(1) \quad (55)$$

This mean contribution will have decayed by a factor  $\exp [(t' - t)/\tau]$  by the time  $t$ . Thus the mean charge at time  $t$  resulting from the accumulation for all previous times is

$$nG'(1) \int_{-\infty}^0 \exp [(t' - t)/\tau] dt' = n\tau G'(1) \quad (56)$$

$G'(1)$  evidently is the mean charge pulse  $\bar{Q}$  in the photomultiplier associated with the entrance of a single gamma ray.

Similarly, the variance in the charge on the condensor at time  $t$  is the integral of the variance of Eq. 54 from  $t' = -\infty$  to  $t' = t$  with a weighting coefficient  $\exp [2(t' - t)/\tau]$  since the decay constant for the square of the charge is twice as large as that for the charge. The result is

$$\frac{n}{2} [G''(1) + G'(1)] \quad (57)$$

The quantity  $G''(1) + G'(1)$  is the mean of the square of the charge pulse associated with a single gamma ray, which we shall designate as  $\bar{Q}^2$ . This is also equal to the variance of the charge pulse associated with a single gamma ray plus  $\bar{Q}^2$ .

The fractional variance of the charge on the condensor is

$$\frac{\sqrt{V}}{M} = \left( \frac{\bar{Q}^2}{2\tau n \bar{Q}^2} \right)^{1/2} \quad (58)$$

The coefficient  $(1/2\tau\eta)^{1/2}$  represents the result obtained by Schiff and Evans for pulses of constant amplitude. The coefficient  $\bar{Q}^2/\bar{Q}^2$  for the thick and thin cases may now be investigated.

#### A. Thick Case

In this case the generating function  $G(\epsilon)$  is  $S_3[G_4[G_5(\epsilon)]]$ . The means and variances of  $G_4$  and  $G_5$  were tabulated in the previous section. The corresponding quantities for  $G_3$  are  $\eta_0 c$  and  $\eta_0^2 c(1-c)$ . By combining the means and variances

$$M = \bar{Q} = 3\rho \eta_0 c \Omega p \quad V = 3\rho^2 \eta_0 c \Omega p [4 + 3\eta_0 \Omega p(1-c)] \quad (59)$$

are obtained. Moreover,

$$\bar{Q}^2 = V + M^2 = 3\rho^2 \eta_0 c \Omega p (4 + 3\eta_0 \Omega p) \quad (60)$$

so that

$$\left(\frac{\bar{Q}^2}{\bar{Q}^2}\right)^{1/2} = \left(\frac{4 + 3\eta_0 \Omega p}{3\eta_0 c \Omega p}\right)^{1/2} \quad (61)$$

As should be expected, this approaches  $1/\sqrt{c}$  when  $\eta_0 \Omega p$  becomes sufficiently large, for the pulses then approach the constant size and the only source of statistical variation is in the random production of luminescent bursts.

#### B. Thin Case

In this case  $G(\epsilon)$  is  $H_3[K_3[G_4[G_5(\epsilon)]]]$  whose averages were tabulated in the previous section.

$$M = \bar{Q} = \frac{3}{2} \rho \alpha \eta_m \Omega p \quad V = 3\rho^2 \alpha \eta_m \Omega p (2 + \eta_m \Omega p) \quad (62)$$

$$Q^2 = 3\rho^2 \alpha \eta_m \Omega p \left[ 2 + \eta_m \Omega p \left( 1 + \frac{3}{4} \alpha \right) \right]$$

$$\left(\frac{Q^2}{\bar{Q}^2}\right)^{1/2} = \left[ \frac{4}{3} \frac{2 + \eta_m \Omega p (1 + \frac{3}{4} \alpha)}{\alpha \eta_m \Omega p} \right]^{1/2}$$

In this case  $Q^2/\bar{Q}^2$  approaches  $(4/3 + \alpha)/\alpha$  when  $\eta_m \Omega p$  becomes sufficiently large.

### 6. CONCLUSIONS

1. The statistical variations in a counting system which consists of a source, a luminescent crystal, and a photomultiplier are examined. It is assumed that the source is constant for a fixed period of time, although it emits particles at random. For definiteness and to provide a maximum degree of statistical variation, it is assumed that the source is a gamma emitter and that only a fraction of the gamma rays fall on the luminescent crystal. The method of generating functions is employed to treat the chain of events which the particles emitted from the source engender. Two oppositely extreme cases are considered, namely, that in which all the energy of a gamma ray which enters the crystal is transferred to the electrons and that in which the gamma ray transfers only a portion of its energy in a manner that depends upon the Compton encounters it makes. The two approximations are referred to as the "thick" and "thin" approximations. The first can be realized by using a crystal which is sufficiently thick that the gamma ray is completely absorbed. The second case can be approximated by using a very thin specimen and using gamma ray energies for which the Compton process predominates.

2. As might be expected the results show that the effectiveness of the system depends upon the ability of the crystal to receive energy from the crystal. They also show that a measure of the effectiveness of the remainder of the system is provided by the quantity

$$x = \eta \Omega p$$

Here  $\eta$  is equal to the number of light quanta,  $\eta$ , produced per gamma ray in the luminescent crystal in the thick case and is,  $\eta_m$ , the maximum number which can be produced per Compton encounter in the thin case.  $\Omega$  is the probability that a light quantum emitted from the crystal will strike the photo-surface of the multiplier, and  $p$  is the probability that a photoelectron will be emitted from the cathode. The system will be a faithful counter of those gamma rays which transfer energy to the crystal provided  $x$  is of the order of 5 or larger. The statistical fluctuations are then determined primarily by the Poisson distribution of encounters in the crystal. If, on the other hand, the current from the multiplier is measured instead of the rate of counts, the contribution of the photomultiplier to the statistical error is appreciable until  $x$  is considerably larger than 5, although this error can be reduced to that corresponding to the Poisson distribution of encounters in the crystal when  $x$  is increased.

The statistics of the case, in which the pulses are fed into a capacitor with a time constant and the voltage of the capacitor is measured, are treated from a standpoint somewhat more general than that considered by Schiff and Evans.

#### REFERENCES

1. J. W. Coltman and F. Marshall, Phys. Rev., 72: 528 (1947); F. Marshall, J. Applied Phys., 18: 512 (1947).
2. I. Broser and H. Kallmann, Z. Naturforsch., 2a: 439 (1937); 642 (1947); I. Broser, L. Herforth, H. Kallmann, and U. Martius, *ibid.*, 3a: 6 (1948).
3. W. Heitler, The Quantum Theory of Radiation, Oxford University Press, 1936.
4. P. W. Engstrom, J. Optical Soc. Am., 37: 420 (1947); G. A. Morton and J. A. Mitchell, R C A Rev., 9: 632 (1948).
5. J. V. Uspensky, Introduction to Mathematical Probability, McGraw-Hill Book Company, Inc., New York, 1937; T. Jorgenson, Am. J. Phys., 16: 285 (1948).
6. L. Herforth and H. Kallmann, Ann. Physik, 4: 231 (1949).
7. J. W. Coltman, E. G. Ebbighausen, and W. Altar, J. Applied Phys., 18: 530 (1947).
8. L. I. Schiff and R. D. Evans, Rev. Sci. Instruments, 7: 456 (1937); L. I. Schiff, Phys. Rev., 50: 88 (1936).

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